

- Variable types: Nominal/Ordinal (Qualitative); Interval/Ratio (Quantitative); Continuous/Discrete
- Binomial:  $p(y) = \binom{n}{y} \pi^y (1 - \pi)^{n-y}$ ,  $y = 0, \dots, n$   
 $E(Y_i) = \pi$ ,  $\text{Var}(Y_i) = \pi(1 - \pi)$   
 $\mu = E(Y) = n\pi$ ,  $\sigma^2 = \text{Var}(Y) = n\pi(1 - \pi)$   
 ML:  $\hat{\pi} = y/n$ ,  $E(\hat{\pi}) = \pi$ ,  $\sigma(\hat{\pi}) = \sqrt{\frac{\pi(1-\pi)}{n}}$   
 Wald:  $z_W = \frac{\hat{\pi} - \pi_0}{\sqrt{\hat{\pi}(1-\hat{\pi})/n}}$ , CI:  $\hat{\pi} \pm z_{\alpha/2} \sqrt{\frac{\hat{\pi}(1-\hat{\pi})}{n}}$   
 Score:  $z_S = \frac{\hat{\pi} - \pi_0}{\sqrt{\pi_0(1-\pi_0)/n}}$
- Multinomial:  $p(n_1, \dots, n_{c-1}) = \left(\frac{n!}{n_1! \dots n_{c-1}!}\right) \pi_1^{n_1} \dots \pi_{c-1}^{n_{c-1}}$   
 $E(n_j) = n\pi_j$ ,  $\text{Var}(n_j) = n\pi_j(1 - \pi_j)$   
 $\text{Cov}(n_j, n_k) = -n\pi_j\pi_k$ , ML estimate:  $\hat{\pi}_j = n_j/n$
- Poisson:  $p(y) = \frac{e^{-\mu} \mu^y}{y!}$ ,  $E(Y) = \text{Var}(Y) = \mu$
- $E(Y) = E[E(Y|\mu)]$ ,  $\text{Var}(Y) = E[\text{Var}(Y|\mu)] + \text{Var}[E(Y|\mu)]$
- Information matrix ( $j, k$ ) element:  $-E\left(\frac{\partial^2 L(\beta)}{\partial \beta_j \partial \beta_k}\right)$
- Standard errors are square roots of diagonal elements of inverse information matrix.
- Information:  $\iota(\theta) = E[-L''(\theta)]$ ,  $\hat{V}(\hat{\theta}) = i^{-1}(\theta)|_{\theta=\hat{\theta}}$
- Wald test stat.  $z = (\hat{\beta} - \beta_0)/SE$ , CI:  $\hat{\beta} \pm z_{\alpha/2} SE$
- Likelihood-ratio test stat.:  $\chi^2_{\dim(H_a \cup H_0) - \dim(H_0)}$   
 $-2 \log(\Lambda) = -2 \log(\ell_0/\ell_1) = -2(L_0 - L_1)$
- Score function:  $u(\beta) = \partial L(\beta)/\partial \beta$
- Chi-squared form of score statistic:  
 $\frac{[u(\beta_0)]^2}{\iota(\beta_0)} = \frac{[\partial L(\beta)/\partial \beta_0]^2}{-E[\partial^2 L(\beta)/\partial \beta_0^2]} = \frac{\text{score}^2}{\text{information}} \sim \chi^2_1$
- Pearson test for specific multinomial  $H_0 : \pi_j = \pi_{j0}$ :  
 $X^2 = \sum_j \frac{(n_j - \mu_j)^2}{\mu_j} \sim \chi^2_{c-1}$ ,  $\mu_j = n\pi_{j0}$
- Likelihood-ratio chi-squared statistic:  
 $G^2 = -2 \log \Lambda = 2 \sum n_j \log(n_j/n\pi_{j0}) \sim \chi^2_{c-1}$

Row	1	2	Marginal
1	$\pi_{11}$ $(\pi_{1 1})$	$\pi_{12}$ $(\pi_{2 1})$	$\pi_{1+} = \sum_j \pi_{1j}$ $(1.0)$
1	$\pi_{21}$ $(\pi_{1 2})$	$\pi_{22}$ $(\pi_{2 2})$	$\pi_{2+} = \sum_j \pi_{2j}$ $(1.0)$
Marg.	$\pi_{+1} = \sum_i \pi_{i1}$	$\pi_{+2} = \sum_i \pi_{i2}$	1.0

- Sensitivity:  $p(\text{test positive} | \text{has disease}) = \pi_{1|1}$
- Specificity:  $p(\text{test negative} | \text{no disease}) = \pi_{2|2}$
- Conditional:  $\pi_{j|i} = \pi_{ij}/\pi_{i+}$  for all  $i$  and  $j$
- Independent:  $\pi_{ij} = \pi_{i+}\pi_{+j}$  for all  $i$  and  $j$
- Homogeneity:  $X$  explanatory,  $\pi_{j|1} = \dots = \pi_{j|I}$
- Difference of proportions:  $\pi_1 - \pi_2$  ( $\pi_i \equiv \pi_{1|i}$ )
- Relative risk:  $\pi_1/\pi_2$  ( $1 \equiv \text{independence}$ )
- Odds:  $\Omega = \pi/(1 - \pi)$  ( $\pi = \Omega/(\Omega + 1)$ )
- Odds ratio:  $\theta = \frac{\Omega_1}{\Omega_2} = \frac{\pi_1/(1 - \pi_1)}{\pi_2/(1 - \pi_2)} = \frac{\pi_{11}\pi_{22}}{\pi_{12}\pi_{21}}$   
 odds ratio = relative risk  $\left(\frac{1-\pi_2}{1-\pi_1}\right)$
- Sample odds ratio:  $\hat{\theta} = n_{11}n_{22}/n_{12}n_{21}$
- Partial table:  $XY$  contingency table for fixed  $Z$
- $XY$  marginal table: sums partial tables over  $Z$
- $2 \times 2 \times K$  tables,  $K \equiv \#$  of control categories:  
 $\{\mu_{ijk}\} \equiv$  cell expected frequencies
- $XY$  conditional odds ratio:  $\theta_{XY(k)} = \frac{\mu_{11k}\mu_{22k}}{\mu_{12k}\mu_{21k}}$
- $XY$  marginal table expected frequencies:  
 $\{\mu_{ij+} = \sum_k \mu_{ijk}\}$
- $XY$  marginal odds ratios:  $\theta_{XY} = \frac{\mu_{11+}\mu_{22+}}{\mu_{12+}\mu_{21+}}$
- Sample counts use cell counts for  $\hat{\theta}_{XY(k)}$  and  $\hat{\theta}_{XY}$ .
- Conditionally independent at level  $k$  of  $Z$ :  
 $P(Y = j|X = i, Z = k) = P(Y = j|Z = k) \forall i, j$
- Conditional independence given  $Z$ :  
 $\pi_{ijk} = \pi_{i+k}\pi_{+jk}/\pi_{++k}$  for all  $i, j, k$
- Homogeneous  $XY$  assoc:  $\theta_{XY(1)} = \dots = \theta_{XY(k)}$
- Mid- $P$ -value:  $\frac{1}{2}P(t = t_o) + P(T > t_o)$

- Sample odds ratio:  $E(\hat{\theta})$  and  $\text{Var}(\hat{\theta})$  undefined.  
 Log odds ratio CI:  $\log \hat{\theta} \pm z_{\alpha/2} \hat{\sigma}(\log \hat{\theta})$   
 $\hat{\sigma}(\log \hat{\theta}) = \left(\frac{1}{n_{11}} + \frac{1}{n_{12}} + \frac{1}{n_{21}} + \frac{1}{n_{22}}\right)^{1/2}$
- Diff. of proportions:  $E(\hat{\pi}_1 - \hat{\pi}_2) = \hat{\pi}_1 - \hat{\pi}_2$   
 Wald CI:  $(\hat{\pi}_1 - \hat{\pi}_2) \pm z_{\alpha/2} \hat{\sigma}(\hat{\pi}_1 - \hat{\pi}_2)$   
 $\hat{\sigma}(\hat{\pi}_1 - \hat{\pi}_2) = \left[\frac{\hat{\pi}_1(1-\hat{\pi}_1)}{n_1} + \frac{\hat{\pi}_2(1-\hat{\pi}_2)}{n_2}\right]^{1/2}$
- Sample relative risk:  $r = \hat{\pi}_1/\hat{\pi}_2$   
 Wald CI:  $\log r \pm z_{\alpha/2} \hat{\sigma}(\log r)$   
 $\sigma(\log r) = \left(\frac{1-\pi_1}{\pi_1 n_1} + \frac{1-\pi_2}{\pi_2 n_2}\right)^{1/2}$
- Delta method:  $T_n \sim N(\theta, \sigma/\sqrt{n})$ ,  $g(\theta)$ :  
 Wald CI:  $g(T_n) \pm z_{\alpha/2} |g'(T_n)| \sigma(T_n) \sqrt{n}$
- Pearson chi-squared test of independence:  
 $X^2 = \sum_i \sum_j \frac{(n_{ij} - \hat{\mu}_{ij})^2}{\hat{\mu}_{ij}} \sim \chi^2_{(I-1)(J-1)}$   
 Est. expected frequencies:  $\hat{\mu}_{ij} = n_{i+}n_{+j}/n$
- Likelihood-ratio chi-squared test:  
 $G^2 = -2 \log \Lambda = 2 \sum_i \sum_j n_{ij} \log(n_{ij}/\hat{\mu}_{ij})$
- Pearson residual:  $e_{ij} = \frac{n_{ij} - \hat{\mu}_{ij}}{\hat{\mu}_{ij}^{1/2}}$
- Standardized Pearson residual:  
 $\frac{n_{ij} - \hat{\mu}_{ij}}{[\hat{\mu}_{ij}(1-p_{+i})(1-p_{+j})]^{1/2}} \sim N(0, 1)$
- $I \times J$  table chi-squared partition:  
 $\frac{\sum_{a < i} \sum_{b < j} n_{ab}}{\sum_{b < j} n_{ib}} \mid \frac{\sum_{a < i} n_{aj}}{n_{ij}}$
- Ordinal:  $M^2 = (n - 1)r^2$
- $\gamma = \frac{\Pi_c - \Pi_d}{\Pi_c + \Pi_d}$ ,  $\hat{\gamma} = \frac{C - D}{C + D}$   
 $\Pi_c = 2 \sum_i \sum_j \pi_{ij} \left(\sum_{h>i} \sum_{k>j} \pi_{hk}\right)$   
 $\Pi_d = 2 \sum_i \sum_j \pi_{ij} \left(\sum_{h>i} \sum_{k<j} \pi_{hk}\right)$
- Fisher's exact test generates all tables consistent with given margin totals:  $H_a : \theta > 1$   
 $p$ -value:  $P(n_{11} \geq t_o)$ ,  $t_o \equiv$  observed  $n_{11}$   
 $p(t) = P(n_{11} = t) = \frac{\binom{n_{1+}}{t} \binom{n_{2+}}{n_{+1}-t}}{\binom{n}{n_{+1}}}$